

Lecture 12 — October 12, 2017

Prof. Jelani Nelson

Scribe: Shyam Narayanan

1 Overview

In the last lecture we discussed efficient algorithms for matrix multiplication and briefly talked about regression. We needed to find an efficient method of generating an ϵ -subspace embedding from last time, since last time our approach required finding the Singular Value Decomposition of A , which is quite slow.

In this lecture we focus on the following:

- Subspace Embeddings
- Regression
- Low-rank approximation

Our general approach is to minimize $\|Ax - b\|$ by looking at $\|\Pi Ax - \Pi b\|$ for some $\Pi \in \mathbb{R}^{d \times n}$, where $d \ll n$.

2 Subspace Embeddings

Recall the following from last lecture:

Definition 1. Π is an ϵ - **subspace embedding** (ϵ -s.e.) for $V = \{x : \exists z \text{ s.t. } x = Uz\}$ (where $U \in \mathbb{R}^{n \times d}$ is some matrix with orthonormal columns, i.e. $U^T U = I$) if

$$\forall x \in V, (1 - \epsilon)\|x\|_2^2 \leq \|\Pi x\|_2^2 \leq (1 + \epsilon)\|x\|_2^2.$$

We showed in the previous lecture that this last condition is equivalent to

$$\|(\Pi U)^T (\Pi U) - I\| \leq \epsilon,$$

where $\|\cdot\|$ represents the operator norm.

We talked about **Singular value decomposition (SVD)**, which tells us for any matrix $A \in \mathbb{R}^{n \times d}$ with rank r , we can write $A = U \Sigma V^T$, where $U \in \mathbb{R}^{n \times r}$, $V \in \mathbb{R}^{d \times r}$, $\Sigma \in \mathbb{R}^{r \times r}$, such that $U^T U = I$, $V^T V = I$, and Σ is a diagonal matrix. If $E = \text{Colspace}(A)$, then letting $\Pi = U^T \in \mathbb{R}^{d \times n}$ gives us $\Pi U = I$ so $\|(\Pi U)^T (\Pi U) - I\| = 0 < \epsilon$. This seems great, but a problem is that solving for U takes time $O(n \cdot d^2)$, which is really slow. So we need to try something different.

We have two ways of constructing subspace embeddings:

1. Sampling
2. “JL” approach

2.1 Sampling

Given as input $A \in \mathbb{R}^{n \times d}$, we want a subspace embedding for $\text{Colspace}(A)$, i.e. $\|\Pi Ax\|_2^2 \approx \|Ax\|_2^2$ for all x . This means we want to preserve $A^T A$, since $\|Ax\|_2^2 = (Ax)^T(Ax) = x^T(A^T A)x$.

Recall that if

$$A = \begin{bmatrix} -a_1^T - \\ \vdots \\ -a_n^T - \end{bmatrix}$$

then

$$A^T A = \sum_{i=1}^n a_i a_i^T.$$

This is a straightforward but valuable fact in linear algebra.

Our goal for constructing Π is to sample each row with some probability p_i . Let

$$\eta_i = \begin{cases} 1 & \text{we keep } a_i \\ 0 & \text{we discard } a_i \end{cases}.$$

Then, we want our matrix

$$\Pi = \begin{bmatrix} \frac{\eta_1}{\sqrt{p_1}} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \frac{\eta_n}{\sqrt{p_n}} \end{bmatrix} \Rightarrow \Pi A = \begin{bmatrix} -\frac{\eta_1}{\sqrt{p_1}} a_1^T - \\ \vdots \\ -\frac{\eta_n}{\sqrt{p_n}} a_n^T - \end{bmatrix}.$$

So this means

$$(\Pi A)^T (\Pi A) = \sum_{i=1}^n \frac{\eta_i}{p_i} a_i a_i^T.$$

Note this means

$$\mathbb{E}[(\Pi A)^T (\Pi A)] = \sum_{i=1}^n \frac{\mathbb{E}[\eta_i]}{p_i} a_i a_i^T = \sum_{i=1}^n a_i a_i^T = A^T A.$$

Note that $\mathbb{E}[\text{number of rows of } A \text{ kept}] = \sum p_i$, so we want to know how small of a p_i we can get away with.

Definition 2. Define

$$R_i = \sup_x \frac{\langle a_i, x \rangle}{\|Ax\|_2^2}.$$

R_i is often thought of as like the “sensitivity” of the row a_i .

Note that $\|Ax\|_2^2 = \sum x^T a_i a_i^T x = \sum \langle a_i, x \rangle^2$.

We want to get some information about p_i given R_i . In fact, we can show the following:

Claim 3. For all i , if $0 < p_i < \frac{R_i}{2}$, then the distribution of Π where we replace $p_i = 0$ is strictly better than the current distribution. In other words, if p_i is not sufficiently large with respect to R_i , it is better that we just set $p_i = 0$.

Proof. Let's fix some i and look at

$$\|\Pi Ax\|_2^2 = \frac{\eta_i}{p_i} \langle a_i, x \rangle^2 + \sum_{j \neq i} \frac{\eta_j}{p_j} \langle a_j, x \rangle^2 \geq \frac{\eta_i}{p_i} \langle a_i, x \rangle^2.$$

Suppose that $p_i \neq 0$. Then, if we were to sample row i (which happens with positive probability),

$$\|\Pi Ax\|_2^2 \geq \frac{1}{p_i} \langle a_i, x \rangle^2$$

for all x . This is true for

$$x^* = \arg \max_x \frac{\langle a_i, x \rangle}{\|Ax\|_2}.$$

But then

$$\|\Pi Ax^*\|_2^2 \geq \frac{R_i}{p_i} \|Ax^*\|_2^2 > 2\|Ax^*\|_2^2,$$

given that $p_i < \frac{R_i}{2}$, which means Π is not ϵ -s.e. Therefore, it is strictly better to let $p_i = 0$ if $p_i < \frac{R_i}{2}$. \square

Definition 4. Given a matrix $M = U\Sigma V^T$ (with $U\Sigma V^T$ as M 's SVD), we define the **pseudoinverse** of M as $M^+ = V\Sigma^{-1}U^T$.

Definition 5. Define $\ell_i = a_i^T (A^T A)^+ a_i$. ℓ_i is called the i th **leverage score** of A .

A lot of papers use leverage score instead of our sensitivity R_i , but it doesn't really matter which one is used. This is because:

Claim 6. $\ell_i = R_i$.

Also, we note the following:

Claim 7. $A(A^T A)^+ A^T$ is the orthogonal projection onto $\text{Colspace}(A)$.

Proof. By looking at the SVD of A , we get

$$A^T A = V\Sigma U^T U\Sigma V^T = V\Sigma^2 V^T.$$

Therefore, $(A^T A)^+ = V\Sigma^{-2} V^T$. This means

$$A(A^T A)^+ A^T = U\Sigma V^T (V\Sigma^{-2} V^T) V\Sigma U^T = U U^T.$$

\square

Note that this implies

$$\ell_i = e_i A(A^T A)^+ A^T e_i = \|U^T e_i\|_2^2 = \|u_i\|^2,$$

where

$$U = \begin{bmatrix} -u_1^T - \\ \vdots \\ -u_n^T - \end{bmatrix}$$

is in $\mathbb{R}^{n \times d}$. Also, if we pick $p_i = \alpha \cdot \ell_i$ for some constant α , then

$$\sum p_i = \alpha \cdot \sum_{i=1}^n \|u_i\|^2 = \alpha \cdot \|U\|_F^2 = \alpha d,$$

since each column of U has unit norm and there are d columns.

It turns out that the following is true:

Theorem 8. [1] *If $p_i \geq \min(1, \alpha \ell_i)$ for all i , and if $\alpha \geq C \cdot \frac{\ln(d/\delta)}{\epsilon^2}$, then*

$$\mathbb{P}(\Pi \text{ is } \epsilon - \text{s.e. for } \text{Colspace}(A)) \geq 1 - \delta$$

Therefore, to compute Π , we just need to compute p_i , but this means we need U , which as we know takes too long to compute. However, there is a fast algorithm that, given A , will compute $\tilde{\ell}_1, \dots, \tilde{\ell}_n$ such that $\forall i, \ell_i \leq \tilde{\ell}_i \leq 2\ell_i$. (Maybe we'll have this on our homework?)

2.2 JL Approach

We will use the technique of “**Oblivious Subspace Embedding**” (OSE) [2].

Definition 9. *A distribution D over $\mathbb{R}^{m \times n}$ is an ϵ, δ -OSE for dimension d if*

$$\forall U \in \mathbb{R}^{n \times d} \text{ s.t. } U^T U = I, \mathbb{P}_{\Pi \sim D}(\|(\Pi U)^T \Pi U - I\| > \epsilon) < \delta.$$

How would we prove that some distribution D is an OSE? There are three main approaches we'll cover:

2.2.1 Nets

We can construct a β -net (in ℓ_2) E' for $E = \{x : x = Uz\}$ for $\beta = \frac{1}{10}$. We can prove that if Π ϵ -preserves all $x \in E'$, then Π ϵ -preserves E . Note that $|E'| = O(\frac{1}{\beta})^d = e^{O(d)}$. Therefore, we need

$$c \cdot \frac{\lg(|E'|/\delta)}{\epsilon^2} = O\left(\frac{d + \lg \frac{1}{\delta}}{\epsilon^2}\right)$$

dimensions, by *JL* lemma.

2.2.2 Moment Method

Let $M = (\Pi U)^T \Pi U - I$. By Markov's inequality, we know that for any $p \geq 1$,

$$\mathbb{P}(\|M\| > \epsilon) < \frac{1}{\epsilon^p} \mathbb{E}(\|M\|^p).$$

Let the eigenvalues of M be $\lambda_1, \dots, \lambda_d$ where $|\lambda_1| \geq |\lambda_2| \geq \dots \geq |\lambda_d|$. Then,

$$\frac{1}{\epsilon^p} \mathbb{E}(\|M\|^p) = \frac{1}{\epsilon^p} \mathbb{E}(\lambda_1^p) \leq \frac{1}{\epsilon^p} \mathbb{E}(\sum \lambda_i^p) = \frac{1}{\epsilon^p} \mathbb{E}(\text{Tr}(M^p)),$$

where we can choose p to be even so λ_i^p is positive. Brute force matrix multiplication tells us that

$$(M^p)_{i,j} = \sum_{i=i_0, i_1, \dots, i_p=j} \prod_{t=0}^{p-1} M_{i_t i_{t+1}}$$

which means that

$$\text{Tr}(M^p) = \sum_{\{i_0, \dots, i_p\}: i_0=i_p} \prod_{t=0}^{p-1} M_{i_t i_{t+1}}.$$

This looks pretty bad, however, it can be useful. As an example, let $p = 2$ and let $\Pi \in \mathbb{R}^{m \times n}$ be the Count Sketch matrix

$$\Pi = \begin{bmatrix} -0- \\ \vdots \\ -\pm 1- \\ \vdots \\ -0- \end{bmatrix}$$

where each column has exactly one nonzero entry. Then, Π is an OSE for $m = \Theta(\frac{d^2}{\epsilon^2 \delta})$ by the moment method for $p = 2$ [3][4][5].

Note that since Π has only one nonzero element per column, $A \mapsto \Pi A$ can be done in time $O(\text{nnz}(A))$, where nnz refers to the **number of nonzero** entries.

The Count Sketch matrix turns out to have the $(\epsilon, \delta, 2) - JL$ moment property for $m = O(\frac{1}{\epsilon^2 \delta})$, which means, as we showed in the previous lecture,

$$\mathbb{P}(\|(\Pi A)^T (\Pi B) - A^T B\|_F > \epsilon \|A\|_F \|B\|_F) < \delta.$$

Now, if $A = B = U$, then $\|A\|_F = \|B\|_F = \sqrt{d}$ so $\|A\|_F \|B\|_F = d$. Letting $\gamma = \frac{\epsilon}{d}$, we need

$$m = \Theta\left(\frac{1}{\gamma^2 \delta}\right) = \Theta\left(\frac{d^2}{\epsilon^2 \delta}\right)$$

rows for the Count Sketch matrix, as mentioned above.

2.2.3 Chaining

We want $\mathbb{E}\|M\| < \epsilon$, where again $M = (\Pi U)^T \Pi U - I$. Recall that $\mathbb{E}\|M\| = \mathbb{E} \sup_{\|x\|_2=1} |x^T M x|$. Then, the following is true:

Theorem 10. [6] Fix $T \subset S^{n-1}$. Then, if $\Pi \in \mathbb{R}^{m \times n}$ with i.i.d. $\mathcal{N}(0, \frac{1}{m})$ entries, then

$$\mathbb{E} \sup_{x \in T} \left| \|\Pi x\|_2^2 - 1 \right| \lesssim \frac{g(T)}{\sqrt{m}} + \frac{g^2(T)}{m},$$

where

$$g(T) = \mathbb{E}_g \sup_{x \in T} \langle g, x \rangle.$$

Now, we can just choose $m \gtrsim \frac{g^2(T)}{\epsilon^2}$ to get the right hand side is $O(\epsilon + \epsilon^2) = O(\epsilon)$.

3 Regression

Recall that we are trying to minimize $\|Ax - b\|$ over x . We try to make faster is to minimize $\|\Pi Ax - \Pi b\|$ where Π has much fewer rows than columns, and where Π is ϵ -s.e. for $\text{span}(b, \text{cols}(A))$ so that $\|\Pi Ax - \Pi b\| \approx \|Ax - b\|$.

Last time, we saw that Π is ϵ -s.e. for $\text{span}(b, \text{cols}(A))$ implies $m = \Theta(d/\epsilon^2)$ is sufficient. We can use fast JL to get an OSE.

We briefly present two other ways:

- The first approach is from [2]. If Π is

1. a $\frac{1}{10}$ -subspace embedding for $\text{Colspace}(A)$ and
2. provides a $\sqrt{\frac{\epsilon}{d}} - \text{AMM}_F$ error for some particular two matrices

then we get some \tilde{x} such that $\|A\tilde{x} - b\|_2^2 \leq (1 + \epsilon) \min \|Ax - b\|_2^2$, and we only need $\frac{d}{\epsilon}$ rows instead of $\frac{d}{\epsilon^2}$ rows.

- The second approach is a gradient descent approach, from [7] [8] [3]. Define $f(x) = \|Ax - b\|_2^2$. Given $x^{(k)}$, we move to $x^{(k+1)} = x^{(k)} - \gamma \nabla f(x_k)$. As long as the ratio of the largest to smallest singular value of A (also called the “condition number” of A or $\kappa(A)$) is not too large, they showed gradient descent converges quickly.

But what if $\kappa(A)$ is not small? Suppose that $\Pi A = U\Sigma V^T, R = V\Sigma^{-1}$. Then, it turns out that $\kappa(AR) = \Theta(1)$, since for all x , $\|\Pi ARx\| = \|Ux\| = \|x\|$, but if Π is ϵ -s.e. for $\text{Colspace}(A)$, then $\|ARx\| \approx \|\Pi ARx\| = \|x\|$, so AR cannot have any eigenvalues that are too small or too large. Therefore, we can do gradient descent with the matrix AR .

References

- [1] Daniel A. Spielman, Nikhil Srivastava. Graph sparsification by effective resistances. *STOC*, 563–568, 2008.
- [2] Tamas Sarlos. Improved Approximation Algorithms for Large Matrices via Random Projections. *FOCS*, 143–152, 2006.
- [3] Kenneth L. Clarkson, David P. Woodruff. Low rank approximation and regression in input sparsity time. *STOC*, 81–90, 2013.
- [4] Jelani Nelson, Huy L. Nguyen. Lower bounds for oblivious subspace embeddings. *CoRR* abs/1308.3280, 2013.
- [5] Xiangrui Meng, Michael W. Mahoney. Low-distortion subspace embeddings in input-sparsity time and applications to robust linear regression. *STOC*, 91–100, 2013.
- [6] Yehoram Gordon. On Milmans inequality and random subspaces which escape through a mesh in \mathbb{R}^n . *Geom. Aspects of Funct. Anal.*, vol. 1317, pages 84–106, 1986-87.

- [7] Vladimir Rokhlin, Mark Tygert. A fast randomized algorithm for overdetermined linear least-squares regression. *PNAS* 105(36): 13212–13217, 2008.
- [8] Haim Avron, Petar Maymounkov, Sivan Toledo. Blendenpik: Supercharging LAPACK’s Least-Squares Solver. *SIAM J. Scientific Computing* 32(3): 1217–1236, 2010.